

Distribution of the GLRT for large data records and weak signal

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EECE 566

Formulation

$$H_0 : \mathbf{x} \sim p(\mathbf{x}; \boldsymbol{\theta}_{r_0}, \boldsymbol{\theta}_{s_0})$$

$$H_1 : \mathbf{x} \sim p(\mathbf{x}; \boldsymbol{\theta}_{r_1}, \boldsymbol{\theta}_{s_1})$$

- $\boldsymbol{\theta}_{r_i}$ parameter of **interest** under H_i
- $\boldsymbol{\theta}_{s_i}$ **nuisance** parameter under H_i

$$(\text{GLRT}): L_G(\mathbf{x}) = \frac{p(\mathbf{x}; \hat{\boldsymbol{\theta}}_{r_1}, \hat{\boldsymbol{\theta}}_{s_1}, H_1)}{p(\mathbf{x}; \hat{\boldsymbol{\theta}}_{r_0}, \hat{\boldsymbol{\theta}}_{s_0}, H_0)} \underset{1}{\overset{0}{\gtrless}} \gamma$$

- $\hat{\boldsymbol{\theta}}_{r_i}, \hat{\boldsymbol{\theta}}_{s_i}$ ML estimates of the parameters under H_i

Theorem (asymptotic distribution of GLRT)

- 1 $\theta_{s_0} = \theta_{s_1}$ nuisance parameters are the same under both hypotheses
- 2 $\|\theta_{r_0} - \theta_{r_1}\|$ small (Weak Signal)
- 3 n large (Large Data Sample, LDR)

$$2 \ln L_G(\mathbf{x}) \stackrel{(a)}{\sim} \begin{cases} \chi_r^2 & \text{under } H_0 \\ \chi_r'^2(\lambda) & \text{under } H_1 \end{cases}$$

λ non-centrality parameter (equivalent of deflection coefficient)
 r dimensionality of θ_{r_0} (the number of parameters of interest)
proof in App. 6A-C

Non-centrality parameter λ

Fisher information matrix

$$[I(\boldsymbol{\theta})]_{i,j} = -E_0 \left[\frac{\partial^2}{\partial \theta_i \partial \theta_j} \ln p(\mathbf{x}; \boldsymbol{\theta}) \right]$$
$$I(\boldsymbol{\theta}_r, \boldsymbol{\theta}_s) = \begin{bmatrix} I_{\theta_r \theta_r}(\boldsymbol{\theta}_r, \boldsymbol{\theta}_r) & I_{\theta_r \theta_s}(\boldsymbol{\theta}_r, \boldsymbol{\theta}_s) \\ I_{\theta_s \theta_r}(\boldsymbol{\theta}_r, \boldsymbol{\theta}_s) & I_{\theta_s \theta_s}(\boldsymbol{\theta}_r, \boldsymbol{\theta}_s) \end{bmatrix}$$

$$\lambda = (\boldsymbol{\theta}_{r_0} - \boldsymbol{\theta}_{r_1})^T I(\lambda) (\boldsymbol{\theta}_{r_0} - \boldsymbol{\theta}_{r_1})$$

$I(\lambda) = I_{\theta_r \theta_r} - I_{\theta_r \theta_s} I_{\theta_s \theta_s}^{-1} I_{\theta_s \theta_r}$ is an $r \times r$ matrix evaluated at the true values $\boldsymbol{\theta}_{r_0}, \boldsymbol{\theta}_{s_0}$ under H_0

Corrolaries

- Pdf under H_0 does not depend on unknown parameters \implies threshold can be determined for a given P_{FA} : Constant False-Alarm detector (CFAR)
- $I_{\theta_r\theta_r} \geq 0$, $I_{\theta_r\theta_s} I_{\theta_s\theta_s}^{-1} I_{\theta_s\theta_r} \geq 0$, both positive semi-definite (PSD) because they are parts of FI matrix, which is PSD \implies the presence of θ_s decreases λ because $\mathbf{x}^T I_{\theta_r\theta_s} I_{\theta_s\theta_s}^{-1} I_{\theta_s\theta_r} \mathbf{x} \geq 0$
 - recall $\lambda = \mathbf{x}^T I(\lambda) \mathbf{x}$, where $\mathbf{x} = \theta_{r_0} - \theta_{r_1}$
 - $I(\lambda) = I_{\theta_r\theta_r} - I_{\theta_r\theta_s} I_{\theta_s\theta_s}^{-1} I_{\theta_s\theta_r} \implies \lambda = \mathbf{x}^T I(\lambda) \mathbf{x} \leq \mathbf{x}^T I_{\theta_r\theta_r} \mathbf{x}$
- If no nuisance parameters, $\theta_s = \{\emptyset\} \implies \theta \equiv \theta_r$ and λ increases, better detection

Example 1: Detection of unknown DC level in WGN (GLRT)

A unknown, σ^2 known, $\theta_{r_1} = \{A\}$, $\theta_{r_0} = \{0\}$, $\theta_{s_1} = \theta_{s_0} = \{\emptyset\}$
From previous lecture on GLRT:

$$2 \ln L_G(\mathbf{x}) = \frac{n\bar{\mathbf{x}}^2}{\sigma^2}$$
$$\bar{\mathbf{x}} \sim \begin{cases} \mathcal{N}(0, \sigma^2/n) & \text{under } H_0 \\ \mathcal{N}(A, \sigma^2/n) & \text{under } H_1 \end{cases}$$

or normalized

$$\frac{\bar{\mathbf{x}}}{\sigma/\sqrt{n}} \sim \begin{cases} \mathcal{N}(0, 1) & \text{under } H_0 \\ \mathcal{N}\left(\frac{A}{\sigma/\sqrt{n}}, 1\right) & \text{under } H_1 \end{cases}$$

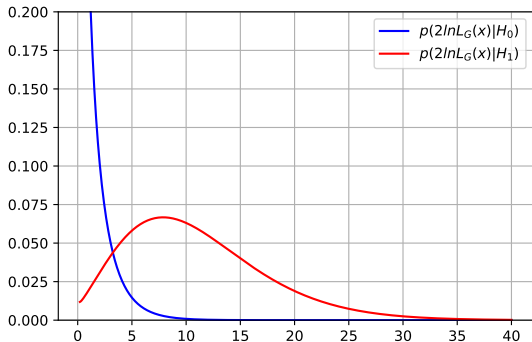
Example 1: Detection of unknown DC level in WGN (GLRT)

$$\left(\frac{\bar{\mathbf{x}}}{\sigma/\sqrt{n}}\right)^2 \sim \begin{cases} \chi_1^2 & \text{under } H_0 \\ \chi_1'^2(\lambda) & \text{under } H_1 \end{cases}$$

- Square of one Gaussian r.v. is χ_1^2 with 1 degree of freedom (pdf is derived in Problem 1 in assignment on statistics)
- The pdf's are **exact**, not just asymptotic in this special case
- Under H_1 , $\chi_1'^2$ is the non-central chi-square with non-centrality $\lambda = \left(\frac{A}{\sigma/\sqrt{n}}\right)^2 = \frac{nA^2}{\sigma^2}$

Example 1: Detection of unknown DC level in WGN (GLRT)

$$n = 10, A = 1, \sigma^2 = 1$$



Example 2: Detection of unknown DC level in WGN of unknown variance (GLRT, LDR)

A unknown, σ^2 unknown, $\theta_{r_1} = \{A\}$, $\theta_{r_0} = \{0\}$,
 $\theta_{s_1} = \theta_{s_0} = \{\sigma^2\}$

From previous lecture on GLRT:

$$2 \ln L_G(\mathbf{x}) = n \ln \left(1 + \frac{\bar{\mathbf{x}}^2}{\hat{\sigma}_1^2} \right)$$
$$2 \ln L_G(\mathbf{x}) \stackrel{(a)}{\sim} \begin{cases} \chi_1^2 & \text{under } H_0 \\ \chi_1'^2(\lambda) & \text{under } H_1 \end{cases}$$

- The distributions are only asymptotically chi-square
- The non-centrality parameter

$$\lambda = A^2 [I_{AA}(0, \sigma^2) - I_{A\sigma^2}(0, \sigma^2)I_{\sigma^2\sigma^2}^{-1}(0, \sigma^2)I_{\sigma^2A}(0, \sigma^2)]$$

Example 2: Detection of **unknown** DC level in WGN of unknown variance (GLRT, LDR)

- Gaussian FI matrix (see slides FI_Gaussian.pdf in Blackboard/Content)

$$I(A, \sigma^2) = \begin{bmatrix} \frac{n}{\sigma^2} & 0 \\ 0 & \frac{n}{2\sigma^4} \end{bmatrix}$$

- Non-centrality λ becomes (since $I_{A\sigma^2}(0, \sigma^2) = I_{\sigma^2 A}(0, \sigma^2) = 0$)

$$\lambda = A^2 [I_{AA}(0, \sigma^2) - I_{A\sigma^2}(0, \sigma^2) I_{\sigma^2 \sigma^2}^{-1}(0, \sigma^2) I_{\sigma^2 A}(0, \sigma^2)]$$

$$\lambda = nA^2/\sigma^2$$

Example 2: Detection of unknown DC level in WGN of unknown variance (GLRT, LDR)

We can set the threshold (we did this before)

$$\begin{aligned} P_{\text{FA}} &= \Pr\{2 \ln L_G(\mathbf{x}) > \gamma | H_0\} \\ P_{\text{FA}} &\stackrel{(a)}{=} \Pr\{\chi_1^2 > \gamma\} \\ &\stackrel{(a)}{=} \Pr\{\mathcal{N}(0, 1)^2 > \gamma\} \\ &\stackrel{(a)}{=} \Pr\{|\mathcal{N}(0, 1)| > \sqrt{\gamma}\} \\ &\stackrel{(a)}{=} \Pr\{\mathcal{N}(0, 1) > \sqrt{\gamma}\} + \Pr\{\mathcal{N}(0, 1) < -\sqrt{\gamma}\} \\ P_{\text{FA}} &\stackrel{(a)}{=} 2\Pr\{\mathcal{N}(0, 1) > \sqrt{\gamma}\} = 2Q(\sqrt{\gamma}) \\ \Rightarrow \gamma &\stackrel{(a)}{=} \left(Q^{-1}(P_{\text{FA}}/2)\right)^2 \end{aligned}$$

Example 2: Detection of unknown DC level in WGN of unknown variance (GLRT, LDR)

And determine the performance (did this before)

$$P_D = \Pr\{2 \ln L_G(\mathbf{x}) > \gamma | H_1\}$$

$$P_D \stackrel{(a)}{=} \Pr\{\chi_1'^2 > \gamma\}$$

$$\stackrel{(a)}{=} \Pr\{\mathcal{N}(\sqrt{\lambda}, 1)^2 > \gamma\}$$

$$\stackrel{(a)}{=} \Pr\{|\mathcal{N}(\sqrt{\lambda}, 1)| > \sqrt{\gamma}\}$$

$$\stackrel{(a)}{=} \Pr\{\mathcal{N}(\sqrt{\lambda}, 1) > \sqrt{\gamma}\} + \Pr\{\mathcal{N}(\sqrt{\lambda}, 1) < -\sqrt{\gamma}\}$$

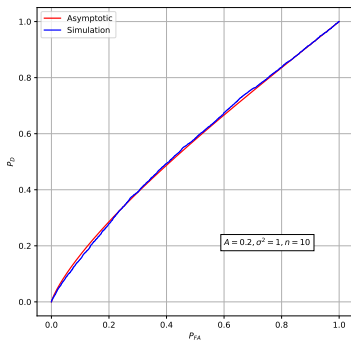
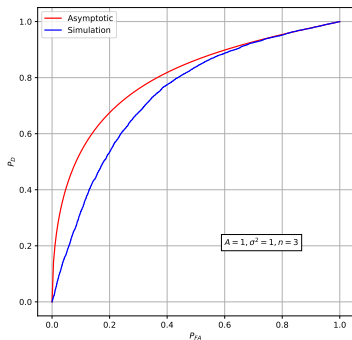
$$\stackrel{(a)}{=} Q(\sqrt{\gamma} - \sqrt{\lambda}) + Q(\sqrt{\gamma} + \sqrt{\lambda})$$

$$P_D \stackrel{(a)}{=} Q\left(Q^{-1}(P_{FA}/2) - \sqrt{\lambda}\right) + Q\left(Q^{-1}(P_{FA}/2) + \sqrt{\lambda}\right)$$

Example 2: Detection of **unknown** DC level in WGN of unknown variance (GLRT, LDR)

- How close is the **actual detector** performance to the **asymptotic approximation**?
 - Monte-Carlo sampling
 - Generate M samples under H_0 and M samples under H_1
 - Use $L_G(\mathbf{x})$ or $T(\mathbf{x}) = \frac{\bar{x}^2}{\hat{\sigma}_1^2}$ (because $\ln(1+x)$ is monotone) and draw the ROC curve
 - Compare to the asymptotic ROC curve derived above
$$P_D = Q\left(Q^{-1}(P_{FA}/2) - \sqrt{\lambda}\right) + Q\left(Q^{-1}(P_{FA}/2) + \sqrt{\lambda}\right)$$

Example 2: Detection of unknown DC level in WGN of unknown variance (GLRT, LDR)



Example 2: Detection of **unknown** DC level in WGN of unknown variance (GLRT, LDR)

